1 How Good Are Symmetric Triangular Synthetic Storms to Represent Real Events

2 for Coastal Hazard Modelling

3 Enrico Duo¹, Marc Sanuy², José A. Jiménez² and Paolo Ciavola¹

⁴ ¹Department of Physics and Earth Science, University of Ferrara, Ferrara, Italy

5 ²Laboratori d'Enginyeria Marítima (LIM), Universitat Politècnica de Catalunya – Barcelona

6 Tech, Barcelona, Spain

7 Corresponding author: E. Duo (<u>duonrc@unife.it</u>).

8

9 Abstract Coastal risk assessments rely on proper quantification of storm-induced erosion and 10 flooding, and often involve calculations via numerical models. When the real time-series data 11 of a storm are not available as forcing conditions and only bulk information is accessible, 12 synthetic simplified time-evolutions are assumed. The most common approach in coastal studies uses a symmetric triangular storm shape, characterised by the assumptions that the peak 13 of the waves occurs in the middle of the storm, and that the forcing varies linearly. This study 14 15 aims to investigate this additional source of uncertainty in hazard estimation, using the XBeach-1D model, to assess the differences in simulated erosion and flooding associated with 16 17 real and synthetic storm definitions. Analysis is performed for real conditions ranging from moderate to extreme at the Northern Adriatic and North-Western Mediterranean coasts, using 18 19 beach profiles ranging from dissipative to reflective. The storm definitions generate 20 considerable differences in terms of wave power and timing at the peak of the storm. When 21 synthetic storms were applied, coastal hazards were not adequately reproduced in most of the simulated cases. The energy of the storms, profile characteristics, local storm climates, and 22 water levels did not consistently influence the differences between the synthetic- and reality-23 based outputs. 24

Keywords Coastal Storm, Mediterranean Sandy Beaches, XBeach, Flooding, Erosion,
Numeric Model Uncertainty

28 1 Introduction

The reliability of the quantification of a hazard component is crucial for coastal risk studies. Coastal inundation and erosion hazards must be satisfactorily evaluated, especially when managing local assessments on sandy beaches. As an example, the magnitudes of a water discharge inundating the hinterland or of an eroded sediment volume are important for adequately evaluating the associated consequences for exposed elements. Moreover, local managers are interested in quantitative information to design risk reduction measures, such as dikes or nourishments, and to prepare management plans.

Nowadays, hazard assessments largely rely on numerical model simulations. Models are indeed 36 capable of reproducing a large amount of processes affecting the interaction between the beach 37 38 morphology and the storm event, to provide results from multiple hazards (Roelvink and Reniers, 2012). Nonetheless, these models rely on assumptions and simplifications that may 39 40 produce unreliable results when compared with observed coastal hazards. As an example, the main factors affecting the simulation of flooding in urbanised coastal areas are linked to the 41 42 mathematical formulations, the topographic data and the forcing boundary conditions (Gallien 43 et al., 2018). Generally, the degree of robustness of a numerical model is related to the data 44 availability and reliability. This is valid for the information on the morphology of the beach, the characteristics of the sediment, and the hydrodynamics. Therefore, the storm event needs 45 46 to be suitably described and included in the numerical models as forcing data. Continuous 47 (observed or hindcasted) storm time-series data of waves and water levels (WLs) are extremely 48 important for capturing the evolution of the event, and thus its dynamic interaction with the beach. 49

When continuous forcing time-series data are unavailable, the event is generally described 50 through observed or assessed bulk information, e.g. maximum significant wave height (Hs), 51 52 peak wave period (Tp), maximum WL (mean sea level+surge+tide), duration (Dur), and main 53 direction (Dir). The lack of continuous data leads to the introduction of simplifications and 54 assumptions to proceed with the analysis of the storm hazard impacts. The most simplified 55 approaches calculate impacts directly using statistical bulk information (see Ranasinghe and Callaghan, 2017). However, accounting for wave and WL variations during the storm is 56 57 necessary for feeding process-based models (see e.g. Roelvink et al., 2009). In these cases, the evolution of the storm must be defined by means of a synthetic shape, hereafter called a 58 synthetic storm (SS), with the assumption that it is representative of the real storm (RS). SSs 59 are regularly used to define the shape of probabilistic storm events (i.e. representative of a 60

61 given return period). A first attempt to standardise a procedure for SS applications can be found in Carley and Cox (2003), wherein they proposed a synthetically-designed storm with 62 exponential-like growth and decay phases, and a symmetrical evolution around the peak. This 63 was obtained by assessing Hs exceedances over various durations and associated with different 64 65 return periods. Simpler approaches have been proposed to adapt synthetic storm shapes to the 66 development of real storms. The triangular shape (Boccotti, 2000; Fedele and Arena, 2009; 67 Corbella and Stretch, 2012a; Laface et al., 2016) is the most frequently applied due to its simplicity, while some studies investigated other shapes, such as the parabola or the trapezoid 68 69 (Martin Soldevilla et al., 2015; Lin Ye et al., 2016), as well as exponential laws (Laface and Arena, 2016). The most recent approaches focused on robust statistical analysis of wave time-70 71 series to model the storm evolution (Solari and Losada, 2018; Lira-Loarca et al., 2020) or to 72 generate joint time-series of wave parameters (Jäger et al., 2019). With some exceptions, most 73 of above-mentioned methods rely on the availability of the storm time-series to properly mimic 74 the storm development. Among all of them, the simplest approach that is widely applied in 75 coastal studies is the symmetric triangular synthetic storm (STSS) (e.g. McCall et al., 2010; Corbella and Stretch, 2012b). It represents the evolution of an event from bulk characteristics 76 77 at the peak and in the storm duration. STSSs are often used to cover all of the possible 78 combinations of forcing (including those not previously recorded) when hazard and risk 79 assessment approaches are applied, by simulating a large number of realistic storm conditions 80 (e.g. Poelhekke et al., 2016; Plomaritis et al., 2018; Sanuy et al., 2018; Santos et al., 2019).

Thus, the use of any type of SS represents a useful approach for coastal hazard assessments, 81 82 and the use of an SS is recommended for planning purposes by Nielsen and Adamantidis (2007). However, SSs show some inherent limitations, and represent an additional source of 83 84 uncertainty in the analyses. Although there are some studies analysing the performance of SSs 85 to represent the storm climate (e.g. Lin-Ye et al., 2016) but only a few analyse their effect in modelling coastal hazards. Sánchez-Arcilla et al. (2009) compared computed erosion impacts 86 from RSs and SSs in the Spanish Mediterranean. The study used schematised, linearly-varying 87 88 Hs and Tp mimicking the shape of the RS, and thus would have had little practical application 89 if only the bulk parameters were known (e.g. as in the case of the STSS). Callaghan et al. (2009) 90 assessed the reliability of an approach proposed by Carley and Cox (2003) for erosion 91 assessments at Narrabeen Beach (Sydney, Australia), by comparing erosion impacts computed 92 from adopting statistical events (i.e. representative of given return periods and simulated with synthetically-designed storms) and statistics of measured impacts. This study found a tendency 93

94 to underestimate computed eroded volumes (EVs) with return periods between three and ten 95 years. However, the results in Callaghan et al. (2009) demonstrate two different components 96 of the uncertainty: the use of the SS, and the uncertainty of the methodology for assigning 97 probabilities to the hazard (e.g. Sanuy et al., 2019). Therefore, the effect of the synthetic approach on the uncertainty was not isolated. More recently, the performance of triangular 98 synthetic storms (including the STSS) has been evaluated for reproducing damage progression 99 100 (Martín-Hidalgo et al., 2014) and overtopping (Martín Soldevilla et al., 2015) in marine 101 structures. Triangular SSs showed a good performance but, depending on the characteristics of 102 the storm, they tend to overestimate or underestimate damage. No study has ever assessed the 103 role of commonly-used SSs in the propagation of uncertainties when modelling both coastal 104 inundation and erosion hazards.

105 Within this context, the main aim of this work is to investigate the differences in storm-induced erosion and inundation assessments associated with the definition of storms (i.e. RS versus SS 106 107 time-series) when using numerical modelling for specific storm conditions. The focus of this 108 study is on the use of the most common and straightforward way of defining a SS, i.e. the STSS. To this end, the magnitude of coastal flooding and erosion is assessed using an extensive 109 dataset of RS data and equivalent synthetic representations. The obtained variations are 110 111 analysed, and are characterised from the differences observed in the storms. The analysis is performed for real conditions typical of the Northern Adriatic and North-Western 112 113 Mediterranean coasts (Figure 1a). These cover beach profiles ranging from dissipative to reflective, and are subjected to storm conditions ranging from moderate to extreme. Storm-114 induced hazards were simulated with the XBeach-1D model (Roelvink et al., 2009). 115



Figure 1. (a) Locations of the sites in the Northern Adriatic and North-Western Mediterranean Seas. The site (f), i.e. Lido Estensi-Spina is located on the (b,d) Emilia-Romagna (Italy) coast; whereas the site (g), i.e. the Tordera Delta is on the coast of (c,e) Catalunya (Spain). The main cities and towns are shown in (d), (e), (f), and (g) as circles. The locations of the wave buoys used to retrieve the wave data used in this study are shown in (d) and (e) as triangles. The partial tracks of the profiles used to select the representative data analysed in this study are shown in (f) and (g) as grey lines.

125 2 Methods and Data

126 2.1 Study area and data

The study area comprises two coastal stretches: in the Northern Adriatic (hereafter NA), Lido 127 degli Estensi-Spina (Italy); and in the North-Western Mediterranean (hereafter NWM), the 128 129 Tordera Delta (Spain) (Figure 1). These two areas are composed by fine and coarse sandy beaches, respectively. Both have been impacted by coastal storms, and they have already been 130 131 classified as critical coastal sectors at the regional level (Armaroli and Duo, 2018; Jiménez et al. 2018). Sun-and-sand tourism is the main coastal economic sector at both sites and, owing 132 133 to this, the related infrastructures and services (e.g. beach facilities, campsites, restaurants) are directly located on the beach, or in the immediate first part of the hinterland. Thus, these 134 135 beaches provide space for accommodating beach users during the bathing season, and protection to the hinterland during the storm season. The general characteristics for each site, 136 137 as well as the main references regarding site conditions, can be found in Table 1. The main 138 data used in the analysis are shown in Table 2.

Site	Site	Sea Basin	Environment	Tidal	Storm	Waves	Main
ID				range	Surge		references
							for regional
							and local
							scales
Italy	Lido degli	Northern	Micro-tidal;	neap:	1-in-	Mean	Masina and
(IT)	Estensi-	Adriatic	Low-	0.3–0.4	10 yrs:	wave	Ciavola,
	Spina	(NA)	energetic;	m	0.72 m	height:	2011;
	(Comacchio,		Dissipative.	spring:	0.72 m	~0.4 m [·]	Armaroli et
	Italy)			0.8–0.9		0.1111,	al., 2012;
				m		Max	Armaroli
						wave	and Duo,
						height:	2018;
						4.6 m*.	Duo et al.,
							2018;
							Sanuy et
							al., 2018.
Spain	Tordera	North-	Micro-tidal;	neap:	1-in-	Mean	Mendoza et
(ES)	Delta	Western	Medium-	0.2–0.25	10 yrs:	wave	al., 2011;
	(Blanes-	Mediterranean	energetic;	m	0.51 m	height:	Jiménez et
	Maresme,	(NWM)	Intermediate-	spring:		~0.7 m:	al., 2018;
	Spain)		reflective.	0.3–0.4		,	Sanuy et
				m		Max	al., 2018;
						wave	Sanuv et
						height:	al., 2019.
						5.4 m**	.,

141 Table 1. General characteristics of study sites.

142 *recorded in February 2015 at the buoy in Figure 1f; **at the virtual node in Figure 1g.

Site	Dataset	Туре	Resolution	Period	Source
ID					
IT	Wave time-	Offshore buoy	0.5 h	2007–	ARPA E-R
	series	Wave buoy at 10 m		2018	Available at:
	(Hs, peak	depth		(83%	https://simc.arpae.it/dext3r/
	wave period	(see Figure 1f)		coverage)	
	(Tp), main				
	direction				
	(Dir))				
	Topography	Lidar	1 × 1 m	October	National Oil Company, Eni
	digital			2014	
	surface				
	model				
	(DSM)				
	Nearshore	Lidar	1 × 1 m	2012	National Oil Company, Eni
	Bathymetry				
	Offshore	Multibeam	1 × 1 m	2013	National Oil Company, Eni
	Bathymetry				
ES	Wave time-	DOW hindcast	1 h	1960–	IH-Cantabria (Reguero et
	serie	20 m depth virtual		2014	al., 2012; Camus et al.,
	(Hs, Tp, Dir)	buoy			2013)
		(see Figure 1g)			
	Bathymetry	Multibeam	$1 \times 1 \text{ m}$	2010	Spanish Ministry of
					Agriculture, Food and
					Environment
	Topography	Lidar	1 × 1 m	2010	Institut Cartogràfic de
	(DSM)				Catalunya
					Available at: <u>www.icgc.cat</u>

145 Table 2. Summary information on the topo-bathymetric and wave datasets.

146

147 2.2 Real storms

148 The first step in the analysis consists of defining the storms. To this end, storms were identified 149 at each site by applying the peak-over-threshold (POT) method, with a double threshold for

150 Hs, i.e. the 0.98 and 0.995 quantiles of the respective time-series, and by imposing a minimum

151 Dur based on local experience (see Table 3). The first Hs threshold (0.98 quantile) was used to calculate Dur, and to define the period between consecutive events. Events with shorter 152 durations than the minimum Dur were not considered. Consecutive peaks with conditions under 153 the threshold lasting less than the meteorological independence criterion (Table 1) were 154 considered as part of the same storm event. The second Hs threshold was applied to identify 155 the most significant storms, which are defined here as extreme events. Table 3 summarises 156 main characteristics of the POT analysis for both sites. A total of 227 storms were identified to 157 build the storm dataset (48 and 179 for the NA and NWM basins, respectively). As both wave 158 159 datasets correspond to different water depths (10 m at NA, deep waters at NWM; see Table 2), the NA storms were linearly back-propagated to the deep waters to generate a consistent 160 161 dataset.

Once the storms were identified, each storm was characterised through a set of wave parameters: Hs at the storm peak (Hs,max); Tp; Dir, and Dur. Then, the energy content (E) of the storm was calculated in the form of a proxy, as previously done by Mendoza et al. (2011), as:

166

$$E = \int Hs^2 dt \tag{1}$$

where t is time in hours. Additionally, the wave power of the storm (P) was calculated to characterize its strength, since induced hazards depend on the rate at which wave energy is delivered (e.g. Burgint et al., 2017), and due to this it is becoming a main parameter to analyse temporal and spatial patterns in storminess (see e.g. Bromisrki and Cayan, 2015). P was calculated as:

172
$$P = \frac{\rho g}{8} \int Hs^2 \cdot Cg \, dt, \tag{2}$$

Where t is time in seconds, ρ is the water density, g is the gravity, and Cg is the group velocity, which depends on Tp and water depth. Since storm definition is specified at the XBeach model outer boundary, and this is located at 20 m water depth, P was calculated by using the intermediate water version. Note that, in this study, E and P are calculated by integration over time for the entire duration of the storm, thus they units are [m²·s] and [(W/m)·s], respectively.

178

Table 3 Characteristics of the peak-over-threshold (POT) analysis for identifying storms at each study site.

Site	Sea Basin	Hs	Hs	Minimum	Meteorological	Nr. of
ID		98%	99.5%	storm	independence	storms
		quantile	quantile	duration	criterion	
IT	NA	1.85 m	2.6 m	4 h	12 h	48
ES	NWM	2 m	2.6 m	6 h	72 h	179

182 2.3 Synthetic storms

To define a SS representing a real event, a simple shape describing the evolution of wave parameters during the storm must be selected (McCall, 2010; Poelhekke et al., 2016; Sanuy et al., 2018). As previously mentioned, this work focuses on the use of STSSs, where Hs linearly grows from the threshold value up to a Hs,max halfway through the storm duration. From there, it linearly decreases down to the threshold value (Figure 2).

To fully define the storm, it is necessary to assign a Tp to each Hs condition. This is a common 188 189 problem in extreme wave analysis, when is necessary to associate a Tp with a height of a given 190 return period (Mathiesen et al 1994). This is a site-specific problem which is solved by deriving 191 empirical relationships, with copula-based approaches being widely used when real storm data 192 are available (e.g. Corbella and Stretch, 2013). However, unless copula-based transformations 193 for any site become available, the most usual way to do it is by using Tp-Hs deterministic 194 relationships which are supplied together with extreme distributions of Hs (see e.g. Sanuy et 195 al. 2019). For instance, in Spain, the State Ports Authority (Puertos del Estado) following Mathiesen et al (1994) provides a specific Tp-Hs relationship to be used together with the 196 197 extreme wave height distribution for different areas along the Spanish coast. It is out of the 198 scope of this work to analyse which is the best way to derive such relationships, thus, site 199 specific relationships were applied. To assign the corresponding wave periods to each STSS, 200 an empirically-derived Tp-Hs linear relationship, separately assessed for each storm dataset 201 (Table 2) by using Hs and Tp bulk data at the peak of the events (for NA: 202 $Tp[s]=1.32 \cdot Hs[m]+3.86$; for the NWM: $Tp[s]=1.75 \cdot Hs[m]+3.69$), is used (see e.g. Mathiesen et al., 1994). The linear fitting resulted in RMSEs ~0.9 s for both datasets, only considering the 203 204 storm peaks. When evaluated for the entire timeseries (i.e., using the real Hs to model Tp), the RMSEs increased to ~1.15 s. With this, the synthetic wave period time-series will depend on 205 the obtained empirical relation Tp-Hs, and on the adopted symmetric triangular shape of the 206 synthetic Hs. Dir would correspond to the mean wave direction during the peak of the event, 207 208 although in this study it is not considered. This is because in this analysis, the worst-case

scenario is investigated, which corresponds to normal incidence. Since this study focuses on the schematization of the wave component, the effects of time-varying WLs (i.e. mean sea level+surges+tides) are ignored, and the WL is assumed to be constant for the duration of the storms.

To compare the SSs and RSs, a set of parameters have been selected. These parameters essentially characterise differences in storm shape (storm peak), E, Tp and P (see Table 4). The

215 peak delay (PD) is defined as the time lag between the peaks of the RSs and SSs (Figure 2).

216



217



219

Table 4. Indicators to compare real and synthetic storms. Subscripts refer to real (r) and synthetic (s) storms.

Symbol	Name			Formula
ΔΕ	Storm	energy	relative	$100 \cdot (E_s - E_r)/E_r$
	differen	ce		
ΔΤρ	Peak	period	relative	$100 \cdot (Tp_s - Tp_r)/Tp_r$
	differen	ce		
ΔPD	Relative	peak dela	ıy	$100 \cdot [t(Hs,max_s) - t(Hs,max_r)]/(0.5 \cdot Dur)$

ΔΡ	Storm wave power relative	$100 \cdot (P_s - P_r)/P_r$
	difference	

223 2.4 Modelling of storm-induced hazards

224 To simulate storm-induced hazards, the process-based morphodynamic model XBeach 225 (Roelvink et al., 2009) was used. It can be considered as a state-of-the-art model for simulating 226 the impact of extreme events, and it is one of the most-used models for this purpose (e.g. 227 McCall et al., 2010; Vousdoukas et al., 2012; Williams et al., 2015; Harley et al., 2016; Passeri et al., 2018). The model was applied in profile mode (1D), similarly to Vousdoukas et al. (2012) 228 and Harley et al. (2016). Beach morphology, WL, waves, and water discharge were simulated 229 230 and stored during the entire simulation of the storms. The parameters of the model were defined as the default values, except for morfac (5), D50, D90 (see Table 5), and bedfriction (white-231 232 colebrook-grainsize). In this way, the friction was calculated as a direct function of the 233 sediment grain size.

234 In this application, topographic and bathymetric datasets (Table 2) of each site were merged to 235 build a coastal digital terrain model, from which a significant number of profiles (i.e. 80 at the 236 site in Italy (IT), Figure 1f; 67 at the site in Spain (ES), Figure 1g) were extracted to describe 237 the local morphology of the beach in detail. At each site, the extracted profiles were classified into five groups, covering a range of local beach morphology. Grouping was performed by 238 239 minimising the variability of all profiles with respect to an average profile, which was used to 240 represent the beach morphology of the sector. This resulted in five average profiles for each 241 site (Figure 3). All the profiles were artificially extended to a 20 m depth for consistency with the forcing time-series. The basic characteristics of the representative profiles and sediments 242 243 (D50 and D90) are summarised in Table 5.

The storm conditions for the simulation consisted of 227 real events (see Section 2.2), and their 245 227 synthetic representations (see Section 2.3). Each real and synthetic event was simulated 246 for each of the 10 profiles. To include the potential variability owing to the mean sea level 247 conditions, three WL scenarios were defined (baseline WL, +0.25 m, +0.75 m). As a result, a 248 total of 13620 simulations were computed.

The obtained results were the morphology and water discharge for each simulation. The water discharge (Q) time-series was extracted for each profile at the locations shown in Figure 3. The

discharge positions were defined in areas that were not significantly affected by erosion for the

entire dataset of simulations, and that were close enough to the shoreline to capture significant







Figure 3. Overview of the profile dataset with indication of the discharge locations.

Site ID	Grain size		Representative	Berm	Slope	Dune	Bar
	[mm]		average	elevation			
	D50	D90	profile				
IT	0.23	0.3	1	1.06 m	0.043	Yes	Yes
			2	0.79 m	0.033	No	Yes
			3	1.00 m	0.031	No	Yes
			4	0.95 m	0.029	No	No
			5	1.11 m	0.005	No	Yes
ES	1.3	1.9	1	3.76 m	0.096	No	No
			2	2.89 m	0.099	No	No
			3	3.11 m	0.117	No	No
			4	2.70 m	0.117	No	No
			5	2.10 m	0.080	No	Yes

258 Table 5. Summary information on the profile dataset.

260 2.5 Analysis of simulated hazards

The EV of the emerged beach (i.e. from the shoreline to where erosion ends) was calculated by comparing the initial and post-storm profiles. The maximum and significant (i.e. the average of the highest third, to capture the average magnitude near the peak of the event) water discharges were calculated (as Qmax and Qs, respectively), as well as the total water volume (TWV) inundating the hinterland. These variables give quantitative information on both the peak of the storm (i.e. Qmax) and its event-integrated values (i.e. EV, Qs, and TWV).

For each variable, the differences between the real- and synthetic-driven outputs were assessed in an event-to-event manner through the expressions shown in Table 6. Positive values of the comparative variables indicate an over-estimation of the STSS in comparison to the RS, and vice versa. The use of relative differences can, however, generate misleading interpretations of the results for high-intensity events, as important absolute differences are smoothed relative to a large hazard output.

273

275	Table 6. Summary of the functions adopted to quantify the comparison between real- and	L
276	synthetic-driven outputs.	

-		
Symbol	Name	Formula
ΔQD	Relative Peak Discharge Delay	$100 \cdot [t(Qmax_s) - t(Qmax_r)]/Dur$
ΔQs	Significant Discharge Relative Difference	$100 \cdot (Qs_s - Qs_r)/Qs_r$
ΔQ _{max}	Maximum Discharge Relative Difference	$100 \cdot (Qmax_s - Qmax_r)/Qmax_r$
ΔTWV	Total Water Volume Relative Difference	$100 \cdot (TWV_s - TWV_r) / TWV_r$
ΔΕV	Eroded Volume Relative Difference	$100 \cdot (EV_s - EV_r) / EV_r$

278

279 3 Results

280 3.1 Storm characteristics

The application of the POT method to both datasets resulted in a total of 227 storms, 48 in the NA, and 179 in the NWM basin. As mentioned before, because the NA wave data were recorded at 10 m depth, the storm Hs values were back-propagated to the deep waters to obtain the corresponding offshore values and thereby generate a consistent dataset. The main characteristics of the identified storms (RS) at each site can be seen in Figure 4.



Figure 4. Main characteristics (wave height (Hs), duration (Dur), energy content (E), and wave power (P)) of RSs at both study sites. Black bars: relative frequency distribution for the whole dataset (NA+NWM); Red bars: relative conditional frequency distribution for Northern Adriatic (NA) storms; Green bars: relative conditional frequency distribution for North-Western Mediterranean (NWM) storms.

293

The comparison between a normalised shape of a RS versus its reproduction by means of the use of SSTS is shown in Figure 5. In addition, the median and associated 75% probability range (given by the 0.175 and 0.825 quantiles) of the normalised Hs time series of both storms are represented. As can be seen, the STSS mimics the typical Hs evolution, although some differences also occur. The average RS shows higher growth rates during a shorter Dur as compared to the average STSS. The average shape of the RS presents a plateau, indicating a natural variability in the occurrence of the peak during the storm. As a difference (and by definition), the average STSS shows a point peak at the middle of the storm. The shadowed areas in Figure 5 represent the variability of the Hs evolution during the storm, which, as expected, is larger for the RS.



304

Figure 5. Significant Hs evolution for representative RSs and STSSs. From left to right: Northern Adriatic (NA) storms; North-Western Mediterranean (NWM) storms; and the whole dataset of storms (NA + NWM). Variables (Hs and Dur) are normalised. Solid lines correspond to the median for each storm type for the entire dataset, and the shadow area delineates the 0.175 and 0.825 quantiles.

310

Figure 6 illustrates a comparison between parameters defining RSs and SSs, in terms of the 311 relative differences in PD, E, Tp, and P. As can be seen, the timing of the storm PD is 312 reasonably well-captured, with more than 60 % of the total cases having a phase lag shorter 313 314 than 6 h. In general terms, the adopted symmetric shape of the SSTS resulted in peaks slightly 315 more frequently delayed with respect to the RS. However, when this parameter is measured in relative terms (Δ PD), the results indicate that 66% of storms present a phase lag of the peak 316 that is longer than 20% of the Dur (as a reference, this corresponds to a phase lag of ~10 h on 317 a 2-day storm). With respect to the E, approximately 40% of the cases were well-reproduced 318 319 by using the STSS as they presented a relative difference smaller than 5%. The remaining cases presented both higher and lower energy values, with a slight tendency to underestimate E. 320 321 Figure 6 also shows the differences in Tp between the STSS and RS. The relative difference 322 for Tp is shown for the storm peaks only (Peak Δ Tp), and as average over the whole duration 323 of the storms (Mean ΔTp). By definition of the adopted approach, Peak ΔTp represents the difference due to the adoption of the Tp-Hs empirical relations (see Section 2.3) alone. In 324

325 general, the results show that the adopted approach (i.e. empirical linear relation Tp-Hs) reasonably reproduces real wave periods at the storm peak (more than 50% of the cases 326 presented an absolute relative difference lower than 5% in Peak Tp). The remaining cases show 327 a slight tendency towards overestimating Tp at the storm peak. On the other hand, Mean ΔTp , 328 which is calculated considering the whole duration of the storm, represents the combined 329 difference due to the empirical relations and the synthetic storm shape. Results show a tendency 330 of the approach to underestimate the Tp evolution of the storm, although the absolute values 331 of Mean Δ Tp are always lower than 20%, and for large part of the dataset (44%) are lower than 332 333 5%. The wave power presents relative differences (ΔP) lower than 5 % in less than 40 % of the cases. Actually, differences are contained within \pm 20 % in almost 70 % of cases, with a 334 tendency towards underestimation. 335

336



337

Figure 6 Variability in storm properties between STSS and RS according to the selected control
parameters (Figure 2 and Table 4).

340

341 3.2 Storm-induced hazards

The previously-obtained differences in storm definition propagate to differences in hazard estimation. Figure 7 shows examples of model outputs from integrating the results of all of the 344 performed simulations. The median of the position of the post-storm profile and normalised discharge time-series and the associated 75% probability ranges given by the 0.825 and 0.175 345 quantiles for the RS and STSS, respectively, are presented for two profiles of the dataset (one 346 intermediate-reflective and one dissipative). The discharge normalization was implemented 347 considering the average value between the real and synthetic Qmax for each combination 348 storm-profile. The normalised discharges in Figure 7 provide information on how the STSS 349 350 and RS compare in different phases of the storm relative to Dur, and cannot be interpreted to compare discharge peaks. This is because all STSSs have their peak in the centre of the storm, 351 352 whereas only 7% of RSs do.

When assessing results across all profiles, the analysed events induced erosion and inundation 353 hazards covering a large range of values (Figure 8). Thus, approximately 60% of the cases 354 355 induced an inner EV larger than 60 m^3/m (this is equivalent to an average beachface retreat of approximately 30 m, assuming 2 m of beachface height), and more than 10% generated an 356 357 erosion larger than 120 m³/m (this is equivalent to an average beachface retreat of approximately 60 m, assuming 2 m of beachface height). With respect to inundation, more than 358 approximately 25% of the events resulted in a TWV overtopping the beach and larger than 100 359 m^3/m (as reference, this is an average discharge of ~0.001 m^3/s over 24 h of storm). 360

The use of the STSS to represent the RS resulted in a general underestimation of storm-induced 361 EVs (Figure 8), with approximately 20% of the cases underestimating the EV by more than 362 20%. With respect to the inundation hazard, the analysed variables were not properly simulated 363 by using the STSS. As seen in Figure 7, the differences in the flood-related hazards are larger. 364 365 In general, and independently of the beach type, the use of the STSS results in an under-366 prediction of the water discharge during most of the event, except during the central phase of the storm, when the discharge tends to be overestimated. This agrees with the phase lags 367 368 obtained for the peak discharge (Δ QD, Figure 8). Overall, only a few cases resulted in a good reproduction of the maximum and/or significant discharges (Qmax and Qs), or the TWV. 369 Notably, most cases underestimated or overestimated these variables with relative errors larger 370 than 20%, with a higher tendency towards underestimation (Figure 8). 371



373

Figure 7. Real- (red) and synthetic-driven (blue) post-storm profiles calculated for the whole dataset of simulations for a predominantly intermediate-reflective (top-left) and dissipative (top-right) beach profile. Real- (red) and synthetic-driven (blue) normalised discharge timeseries calculated for the whole dataset of simulations for a predominantly intermediatereflective (bottom-left) and dissipative (bottom-right) beach profile. All graphs are represented by the median (solid line) and the 75% of the dataset (shaded area) given by the 0.175 and 0.825 quantiles.



Figure 8. Variability in storm-induced hazards between STSS and RS according to the selectedcontrol parameters (Table 6).

385

386 4 Discussion

The analysis has shown that, although using synthetic time-series to represent wave forcing for simulating storm-induced coastal hazards is a widely-used approach (e.g. McCall et al., 2010; Corbella and Stretch, 2012b; Poelhekke et al., 2016; Plomaritis et al., 2018; Sanuy et al., 2018), the obtained results can significantly differ than those obtained using the real time-series they are intended to represent. This study represents the first attempt to quantify the uncertainty related to the use of these types of synthetic events in deterministic modelling.

393 The use of an STSS can be discussed in two different and complementary ways. The first one 394 regards how well this approach represents the characteristics of an RS. The obtained results showed that, for this purpose, the use of an STSS provides a reasonable representation of 395 reality, as it implies a perfect representation of Hs at Hs, max and the Dur of RSs. When the 396 397 adopted shape has a potential influence on the magnitude of a variable to be characterised, the results begin to differ (e.g. E content). Thus, the selected triangular shape determines the PD 398 399 between both approaches. As has been shown here, even when the analysed storms are 400 retrieved from localised areas (two in this case) where the meteorology presents well-defined

401 and stable patterns, the peak occurs at different phases of the RS development, depending on specific conditions. This results in a relatively wide fraction of the storm duration where the 402 403 peaks can be verified, as contrasted with a single fixed point in the STSS. This prevents the proper representation of the storm growth and relaxation phases and, in consequence, 404 405 potentially affects any process depending on these characteristics. Regarding wave periods, the 406 adopted linear fitting approach (i.e., Tp-Hs) introduces additional errors. The RMSEs evaluated 407 for the linear fit are low (~ 0.9 - 1.15 s; see Section 2.3) for both wave datasets and, as a consequence, the Tp at the storm peaks is reasonably well-captured (Peak Δ Tp) because they 408 409 are only affected by the uncertainty of the adopted linear model. However, because Tp values within the synthetic storm depend on the adopted storm shape - triangular in this case (see 410 Section 2.3) – the reproduction of Tp during the entire storm was less accurate (Mean Δ Tp). 411 412 Since P depends on Hs (thus, on E) and Tp, errors in both variables are transferred to errors in P. 413

414 The second consideration regards how changes in storm properties are transferred to storminduced hazards. As opposed to the previous parameters, according to the obtained results, the 415 adopted STSS has important effects on the reproduction of the induced hazards. Indeed, the 416 storm-induced erosion was properly captured in just 22% of cases, whereas the TWV 417 inundating the hinterland was properly captured in only 4% of cases. The better representation 418 419 of the erosion hazard is a consequence of the morphodynamic feedback taking place during the 420 impact of the storm, where the modifications of the beach morphology affect beach overtopping. In consequence, errors in beach morphology reproduction will propagate (and 421 expand) to beach inundation. 422

423 In this study, the differences in the EV (Δ EV) were strongly related with differences in P (Δ P) between the real and triangular time-series (Figure 9). Secondarily, ΔEV are related to 424 425 differences in wave period (ΔTp), in storm energy (ΔE), and the delay of the peak (ΔPD). As expected, consistent under- and over-estimation of the wave power lead to under- and over-426 427 estimation of the eroded volume. Thus, if the wave power is not well represented, models 428 based on the average equations of mass and momentum cannot properly compute erosion and 429 flooding. In this sense, it has to be considered that most of current definition of synthetic storms, and STSS in particular, are based on representing wave height and, in consequence, 430 431 they do not necessarily conserve wave power during the storm. Notably, approaches to design SS based on P conservation may solve this issue, but they are not applicable when the only 432 available information are the bulk characteristics (at the peak) of the storm event. In Figure 9, 433

 ΔE and Peak ΔTp are only moderately linked to ΔEV when considered separately, while the dependency is emphasized when considering Mean ΔTp . This suggests that both the initial assumptions, on the adopted synthetic shape and on the Tp-Hs relation, affect the proper assessment of EV. However the contribution of the adopted shape has a double impact as directly affecting E (i.e. through Hs) and the Tp time-series of the storm. Note that, the under-/over-estimation of the EV was also linked to the delay of the peak (ΔPD >20%) and storm peak anticipation (ΔPD <-20%) respectively.

441



442

Figure 9. Relation between the eroded volume relative difference (ΔEV) with the variables 443 describing the relative differences between the real and triangular time-series. The relative 444 conditional frequency distributions are shown through coloured tables, where each row 445 446 represents a conditioning range of ΔEV . On the left, top and bottom: distribution of eroded volume relative difference (ΔEV) for the whole dataset. From left to right, top to bottom: 447 448 conditional distributions of energy relative difference (ΔE); period mean relative difference (Mean ΔTp); relative peak delay (ΔPD); period at the peak of the storm relative difference 449 450 (ΔTp) ; and wave power relative difference (ΔP) .

452 Sánchez-Arcilla et al. (2009) also compared the use of RSs and SSs to assess beach erosion using the Sbeach model (Larson and Kraus, 1989). In their study, they used simplified Hs and 453 454 Tp time-series in linear segments following the evolution of the RSs, and thereby captured storm peaks. Their study showed an over-estimation of EVs and shoreline erosion when using 455 a synthetic event, possibly owing to the fact that the approach over-estimated E and Tp, in 456 457 general. The present study, however, evidenced a general under-estimation of the EV as shown in Figure 8. This was linked (Figure 9) to the more frequent under-estimation of P with the 458 STSS, which is determined by under-estimation of E and Tp. Such differences between both 459 460 studies reflect the use of a different approach to represent the storm evolution. Despite this, the differences between real- and synthetic-based outputs were smaller in Sánchez-Arcilla et al 461 (2009) than those found in this study. However, to apply that approach, the shape of the event 462 must be known *a priori* to mimic the storm evolution, whereas the STSS approach only requires 463 storm bulk information. In addition to this, the number of cases simulated here to obtain a 464 robust statistic of errors is much larger, and covers a wider range of conditions than in Sánchez-465 Arcilla et al (2009). 466

The apparent trend highlighted for the ΔEV - ΔE and ΔEV - ΔTp relationships, agrees with the 467 468 findings of McCall et al. (2010). In that work, the authors performed a sensitivity analysis of a 469 2D XBeach model of the barrier island of Santa Rosa (FL, US), varying the synthetic symmetric triangular Hs, and the Tp time-series of the Hurricane Ivan event (the base case) by 470 471 the 30%. Notably, the variation introduced on the wave time-series of the base case did not influence its symmetric triangular shape. An analysis of the morphological impact on a 472 473 foredune area showed that, in addition to expected changes in the EV following changes in Hs (and thus, E), the varying Tp conditions (Mean and Peak $\Delta Tp = \pm 30\%$) resulted, in the under-474 475 estimation ($\Delta EV \sim -30\%$) and over-estimation ($\Delta EV \sim 18\%$) of the EV, respectively. However, 476 the same study also concluded that the erosion model output was more sensitive to (some) sediment transport parameters than to varying hydrodynamic conditions. This suggests that the 477 differences induced using triangular storms (or SSs, in general) can potentially be compensated 478 479 for by a calibration process. However, as the results obtained in this study show both underand over-prediction, deeper investigations are required to verify this hypothesis under a wide 480 481 range of storm conditions.

The obtained results demonstrate the existence of a strong relation between differences in erosion and inundation hazards (see Figure 10). The differences in the EV (Δ EV) and the phase lag of the water discharge (Δ QD) are linked, confirming the importance of morphodynamic

feedback when simulating coastal inundation. A good/reasonable agreement (between real and 485 triangular storms) on the computed EV ($|\Delta EV| < 20\%$) leads to a good agreement on the 486 positioning of the peak of the water discharge ($|\Delta QD| < 5\%$). This should be important when the 487 interest is in accurately timing the peak of the floodwater volume. However, this fine 488 reproduction of the peak timing does not necessarily imply that the total floodwater during the 489 event will be accurately reproduced. In fact, the obtained results showed that a good 490 491 reproduction of the EV ($|\Delta EV| < 5\%$) is not accompanied by a good simulation of the inundation $(|\Delta TWV| < 5\%)$. Despite this, under- or over-estimation of erosion $(|\Delta EV| > 5\%)$ leads to strong 492 493 under- or over-estimation of inundation (|TWV|>20%), respectively.

494



495

Figure 10. Relation between the ΔEV with the relative differences in flooding-related variables. The relative conditional frequency distributions are shown through coloured tables, where each row represents a conditioning range of ΔEV . From left to right: distribution of ΔEV for the whole dataset; conditional distributions of relative peak discharge delay (ΔQD); and total water volume relative difference (ΔTWV).

501

To determine if the previously-observed differences are related to the structure of simulated 502 conditions, they were further analysed according to the energy of the storm, the profile 503 504 conditions (dissipative or reflective), location (storm dataset), and WL (Figure 11). Focusing on the -20% to +20% range of uncertainty in the hazard estimation, the results presented in 505 Section 3 are not strongly conditioned by any of the analysed conditions. Although a slightly 506 507 better estimation of EVs is obtained for reflective conditions and extremely energetic storms, 508 the obtained results are consistently homogeneous throughout the dataset, especially when 509 looking at the relative differences between -5% and +5%.



Figure 11. Relation between ΔEV (on the left) and ΔTWV (on the right) with (from top to bottom) the storm energy (E) class; the profile characteristics (reflective, dissipative); location (i.e. the storm sub-datasets: Northern Adriatic, NA; North-Western Mediterranean, NWM); and water level (WL). The relative conditional frequency distributions of ΔEV and ΔTWV are shown through coloured tables, where each row represents a different conditioning range of the analysed variables.

518

519 5 Conclusions

520 This study investigated the differences generated when simulating the hazard impacts of coastal 521 storms using a STSS of waves, instead of the real data. It was demonstrated that the synthetic 522 method, applied in an event-to-event manner, leads to highly uncertain and misleading 523 deterministic hazard assessments, strongly limiting the reliability of the modelling approach.

After analysing the computed differences in reproducing storm-induced hazards by using STSSs, it can be concluded that they hardly reproduce the real magnitude with independence from the structure of storms or profiles. Differences in wave power are the dominant factor in transferring errors to hazards, and are determined by differences in storm energy (i.e., 528 significant wave height) and period. These are mainly controlled by the adopted synthetic 529 shape, which directly affects the synthetic significant wave height (and thus, energy), and 530 indirectly affects wave period while calculating it with empirical predictive relations (i.e., Tp-531 Hs) applied to the synthetic wave heights; and secondarily, by the empirical predictive relations 522 (i.e., Tp-Uc) which directly affects the synthetic period.

532 (i.e., Tp-Hs), which directly affects the synthetic period.

This is applicable to the range of simulated conditions, and permits one to conclude that although the use of STSSs adequately reproduces some of the main bulk variables defining the storm, they only reasonably reproduce the storm-induced hazard magnitude, i.e. accepting uncertainty in the order of (or greater than) +20% and -20%. Notwithstanding the fact that this type of synthetic approach has been used in recent projects and for engineering purposes, its use should be discouraged, whereas its results should be carefully discussed considering the shortcomings related to its use.

540 This highlights the need for further investigations towards a generalised synthetic approach 541 that can optimise the simulation of coastal hazards, while minimizing the uncertainty related 542 to the use of design events.

543

544 **Conflict of interest statement.** The authors declare that they have no known competing 545 financial interests or personal relationships that could have appeared to influence the work 546 reported in this paper.

547 CRediT authorship contribution statement. Enrico Duo: Conceptualization, Methodology,
548 Software, Formal analysis, Investigation, Data Curation, , Writing - Review & Editing; Marc
549 Sanuy: Conceptualization, Methodology, Software, Formal analysis, Investigation, Data
550 Curation, , Writing - Review & Editing; José A. Jiménez: Conceptualization, Funding
551 acquisition, Writing - Review & Editing, Supervision; Paolo Ciavola: Funding acquisition,
552 Writing - Review & Editing, Supervision.

Acknowledgements. The authors thank: Eni and ARPA-ER for providing topo-bathymetric and wave data for the Italian site; the Institut Cartogràfic i Geológic de Catalunya for supplying topographic data, and IH-Cantabria for supplying wave and water level data for the Spanish site. The authors thank Clara Armaroli, Ap van Dongeren, Tom Spencer, Óscar Ferreira and Diego Vicinanza, for their valuable comments and suggestions at early stages of the manuscript preparation. 559 **Funding.** The work of E. Duo was supported by a PhD grant at the Department of Physics and Earth Science of the University of Ferrara, additional funding from the contribution "5 per 560 561 mille assegnato all'Università di Ferrara - dichiarazione dei redditi dell'anno 2013" assigned through the "Bando Giovani Ricercatori 2016" of the University of Ferrara, and the EU H2020 562 ANYWHERE (GA 700099; www.anywhere-h2020.eu). The work of M. Sanuy and J.A. 563 Jiménez has been done in the framework of the M-CostAdapt (CTM2017-83655-C2-1-R) 564 565 research project (MINECO/AEI/FEDER, UE). Marc Sanuy was supported by a PhD grant from the Spanish Ministry of Education, Culture and Sport. 566

567

568 References

- Armaroli, C. and Duo, E., 2018. Validation of the coastal storm risk assessment framework
 along the Emilia-Romagna coast. Coast. Eng. 134, 159–167.
 https://doi.org/10.1016/j.coastaleng.2017.08.014
- 572 Armaroli, C., Ciavola, P., Perini, L., Calabrese, L., Lorito, S., Valentini, A., and Masina, M.,
- 573 2012. Critical storm thresholds for significant morphological changes and damage along the
- 574 Emilia-Romagna coastline, Italy, Geomorphology, 143–144, 34–51.
- 575 https://doi.org/10.1016/j.geomorph.2011.09.006
- 576 Boccotti, P., 2000. Wave Mechanics for Ocean Engineering. Elsevier Science, Oxford.
- 577 Bromirski, P.D., and Cayan, D.R. (2015). Wave power variability and trends across the N orth
- Atlantic influenced by decadal climate patterns. J. Geophys. Res. Oceans. 120(5), 3419-3443.
 https://doi.org/10.1002/2014JC010440
- 580 Burvingt, O., Masselink, G., Russell, P., Scott, T., 2017. Classification of beach response to
- 581
 extreme
 storms.
 Geomorphology,
 295,
 722-737.

 582
 https://doi.org/10.1016/j.geomorph.2017.07.022
 10.1016/j.geomorph.2017.07.022
 10.1016/j.geomorph.2017.07.022
- Callaghan, D.P., Ranasinghe, R., Short, A., 2009. Quantifying the storm erosion hazard for
 coastal planning. Coast. Eng. 56, 90–93. https://doi.org/10.1016/j.coastaleng.2008.10.003
- 585 Camus, P., Mendez, F.J., Medina, R., Tomas, A., and Izaguirre, C., 2013. High resolution
- 586 downscaled ocean waves (DOW) reanalysis in coastal areas. Coast. Eng. 72, 56-68.
- 587 <u>https://doi.org/10.1016/j.coastaleng.2012.09.002</u>

- Carley, J.T., Cox, R.J., 2003. A methodology for utilising time-dependent beach erosion
 models for design events, in: Kench, P.S., Hume, T.M. (Eds.), Coasts & Ports 2003
 Australasian Conference: Proceedings of the 16th Australasian Coastal and Ocean Engineering
 Conference, the 9th Australasian Port and Harbour Conference and the Annual New Zealand
 Coastal Society Conference, Auckland, New Zealand, 9-1. Institution of Engineers, Australia,
 pp. 587–595.
- Corbella, S. and Stretch, D.D., 2012a. Multivariate return periods of sea storms for coastal
 erosion risk assessment. Nat. Hazards Earth Syst. Sci. 12, 2699–2708.
 https://doi.org/10.5194/nhess-12-2699-2012
- 597 Corbella, S., and Stretch, D.D., 2012b. Predicting coastal erosion trends using non-stationary
 598 statistics and process-based models. Coast. Eng. 70, 40–49.
 599 <u>https://doi.org/10.1016/j.coastaleng.2012.06.004</u>
- Corbella, S., and Stretch, D.D., 2013. Simulating a multivariate sea storm using Archimedean
 copulas. Coast. Eng. .76, 68-78. https://doi.org/10.1016/j.coastaleng.2013.01.011
- Duo, E., Trembanis, A. C., Dohner, S., Grottoli, E., and Ciavola, P., 2018. Local-scale postevent assessments with GPS and UAV-based quick-response surveys: a pilot case from the
 Emilia–Romagna (Italy) coast. Nat. Hazards Earth Syst. Sci. 18, 2969-2989.
 https://doi.org/10.5194/nhess-18-2969-2018
- Fedele, F., and Arena, F., 2009. The equivalent power storm model for long-term predictions
 of extreme wave events. Proc. of the 28th International Conference on Ocean, Offshore and
 Arctic Engineering (OMAE 2009). American Society of Mechanical Engineers (ASME).
 https://doi.org/10.1115/OMAE2009-79597
- Gallien, T.W., Kalligeris, N., Delisle, M.-P.C., Tang, B.-X., Lucey, J.T.D., Winters, M.A.,
- 611 2018. Coastal flood modeling challenges in defended urban backshores. Geosciences, 8, 450.
- 612 https://doi.org/10.3390/geosciences8120450
- Harley, M.D., Valentini, A., Armaroli, C., Perini, L., Calabrese, L., Ciavola, P., 2016. Can an
 early-warning system help minimize the impacts of coastal storms? A case study of the 2012
 Halloween storm, northern Italy. Nat. Hazards Earth Syst. Sci. 16, 209–222.
 https://doi.org/10.5194/nhess-16-209-2016

- Jäger, W.S., Nagler, T., Czado, C., and McCall, R.T., 2019. A statistical simulation method for
- 618 joint time series of non-stationary hourly wave parameters. Coast. Eng. 146, 14-31.
- 619 https://doi.org/10.1016/j.coastaleng.2018.11.003
- 620 Jiménez, J. A., Sanuy, M., Ballesteros, C. and Valdemoro, H. I., 2018. The Tordera Delta, a
- hotspot to storm impacts in the coast northwards of Barcelona (NW Mediterranean). Coast.
- 622 Eng. 134, 148-158. https://doi.org/10.1016/j.coastaleng.2017.08.012
- Laface, V., and Arena, F., 2016. A new equivalent exponential storm model for long-term
 statistics of ocean waves. Coast. Eng. 116, 133-151.
 https://doi.org/10.1016/j.coastaleng.2016.06.011
- Laface, V., Malara, G., Romolo, A., Arena, F., 2016. Peak over threshold vis-à-vis equivalent
- triangular storm: Return value sensitivity to storm threshold. Coast. Eng. 116, 220-235.
 https://doi.org/10.1016/j.coastaleng.2016.06.009
- 629 Larson, M., and Kraus, N.C., 1989. SBEACH: Numerical Model for Simulating Storm-Induced
- 630 Beach Change Report 1: Empirical Foundation and Model Development. Technical Report
- 631 CERC- 89-9, Coastal Engineering Research Center, Water- ways Experiment Station,
- 632 Vicksburg, Mississippi. https://doi.org/10.5962/bhl.title.47893
- 633 Lin-Ye, J., Garcia-Leon, M., Garcia, V., and Sánchez-Arcilla, A., 2016. A multivariate
- 634 statistical model of extreme events: An application to the Catalan coast. Coast. Eng. 117, 138–
- 635 156. https://doi.org/10.1016/j.coastaleng.2016.08.002
- 636 Lira-Loarca, A., Cobos, M., Losada, M.A., and Baquerizo, A., 2020. Storm characterization
- and simulation for damage evolution models of maritime structures. Coast. Eng. 156, 103620.
- 638 https://doi.org/10.1016/j.coastaleng.2019.103620
- 639 Martín-Hidalgo, M., Martín-Soldevilla, M.J., Negro, V., Aberturas, P., and López-Gutiérrez,
- 640 J.S., 2014. Storm evolution characterization for analysing stone armour damage progression.
- 641 Coast. Eng. 85, 1–11. https://doi.org/10.1016/j.coastaleng.2013.11.008
- 642 Martín-Soldevilla, M.J., Martín-Hidalgo, M., Negro, V., López-Gutiérrez, J.S., and Aberturas,
- 643 P., 2015. Improvement of theoretical storm characterization for different climate conditions.
- 644 Coast. Eng. 96, 71–80. https://doi.org/10.1016/j.coastaleng.2014.11.004
- Masina, M., and Ciavola, P., 2011. Analisi dei livelli marini estremi e delle acque alte lungo il
- 646 litorale ravennate. Studi Costieri 18, 87–101.

- 647 Mathiesen, M., Goda, Y., Hawkes, P. J., Mansard, E., Martín, M. J., Peltier, E., Thompson, E.
- F. and Van Vledder, G., 1994. Recommended practice for extreme wave analysis. J. Hydraul.
- 649 Res. 32(6), 803–814. https://doi.org/10.1080/00221689409498691
- 650 McCall, R.T., Van Thiel de Vries, J.S.M., Plant, N.G., Van Dongeren, A.R., Roelvink, J.A.,
- Thompson, D.M., Reniers, A.J.H.M., 2010. Two-dimensional time dependent hurricane
- overwash and erosion modeling at Santa Rosa Island. Coast. Eng. 57, 668-683.
- 653 https://doi.org/10.1016/j.coastaleng.2010.02.006
- Mendoza, E. T., Jimenez, J. A. and Mateo, J., 2011. A coastal storms intensity scale for the
- 655 Catalan sea (NW Mediterranean). Nat. Hazards Earth Syst. Sci. 11, 2453–2462.
 656 https://doi.org/10.5194/nhess-11-2453-2011
- 657 Nielsen, A.F., Adamantidis, C.A., 2007. Defining the Storm Erosion Hazard for Beaches. Aust.
- 658 J. Civ. Eng. 3, 39–50. https://doi.org/10.1080/14488353.2007.11463920
- 659 Passeri, D.L., Bilskie, M.V., Plant, N.G., Long, J.W. and Hagen, S.C., 2018. Dynamic
- 660 modeling of barrier island response to hurricane storm surge under future sea level rise.
- 661 Climatic Change 149(3-4), 413-425. https://doi.org/10.1007/s10584-018-2245-8
- 662 Plomaritis, T.A., Costas, S., Ferreira, Ó., 2017. Use of a Bayesian Network for coastal hazards,
- 663 impact and disaster risk reduction assessment at a coastal barrier (Ria Formosa, Portugal).
- 664 Coast. Eng. 134, 134-147. https://doi.org/10.1016/j.coastaleng.2017.07.003
- Poelhekke, L., Jäger, W.S., van Dongeren, A., Plomaritis, T.A., McCall, R., Ferreira, Ó., 2016.
- Predicting coastal hazards for sandy coasts with a Bayesian Network. Coast. Eng. 118, 21–34.
 https://doi.org/10.1016/j.coastaleng.2016.08.011
- 668 Ranasinghe, R., Callaghan, D., 2017. Assessing Storm Erosion Hazards, in: Ciavola, P., Coco,
- 669 G. (Eds.), Coastal Storms: Processes and Impacts. John Wiley & Sons Ltd., pp. 241–256.
- 670 Reguero, B.G., Menéndez, M., Méndez, F.J., Mínguez, R. and Losada, I.J., 2012. A Global
- 671 Ocean Wave (GOW) calibrated reanalysis from 1948 onwards. Coast. Eng. 65, 38-55.
- 672 https://doi.org/10.1016/j.coastaleng.2012.03.003
- 673 Roelvink, D., Reniers, A., 2012. A guide to modeling coastal morphology. World Scientific
- 674 Publishing Co. Pte. Ltd. https://doi.org/10.1142/9789814304269

- 675 Roelvink, D., Reniers, A., van Dongeren, A., van Thiel de Vries, J., McCall, R., Lescinski, J.,
- 676 2009. Modelling storm impacts on beaches, dunes and barrier islands. Coast. Eng. 56, 1133–
- 677 1152. https://doi.org/10.1016/j.coastaleng.2009.08.006
- 678 Sánchez-Arcilla, A., Mendoza, E.T., Jiménez, J.A., Peña, C., Galofré, J., Novoa, M., 2009.

679 Beach Erosion and Storm Parameters: Uncertainties for the Spanish Mediterranean., in: McKee

- 680 Smith, J. (Ed.), Coastal Engineering 2008 Proceedings of the 31st International Conference
- 681 Hamburg, Germany, 31 August 5 September 2008. World Scientific, pp. 2352–2362.
- 682 https://doi.org/10.1142/9789814277426_0194
- 683 Santos, V.M., Wahl, T., Long, J.W., Passeri, D.L., and Plant, N. G., (2019). Combining
- numerical and statistical models to predict storm-induced dune erosion. J. Geophys. Res. Earth

685 Surf. 124, 1817–1834. https://doi.org/10.1029/2019JF005016

- 686 Sanuy, M., Duo, E., Jäger, W.S., Ciavola, P., Jiménez, J.A., 2018. Linking source with
- consequences of coastal storm impacts for climate change and risk reduction scenarios for
 Mediterranean sandy beaches. Nat. Hazards Earth Syst. Sci. 18, 1825–1847.
 <u>https://doi.org/10.5194/nhess-18-1825-2018</u>
- 690 Sanuy M., Jiménez J.A., Ortego M.I., Toimil A., 2019. Differences in assigning probabilities
- 691 to coastal inundation hazard estimators: Event versus response approaches. J. Flood Risk
- 692 Management. e12557. <u>https://doi.org/10.1111/jfr3.12557</u>
- Solari, S., and Losada, M.A., 2018. Simulation of sea storms including multivariate storm
 evolution. Proc. 36th Int. Conf. on Coast. Eng. ASCE, papers.35.
 https://doi.org/10.9753/icce.v36.papers.35
- Vousdoukas, M.I., Ferreira, Ó., Almeida, L.P. and Pacheco, A., 2012. Toward reliable stormhazard forecasts: XBeach calibration and its potential application in an operational earlywarning system. Ocean Dynamics 62(7), 1001-1015. https://doi.org/10.1007/s10236-0120544-6
- Williams, J.J., Esteves, L.S. and Rochford, L.A., 2015. Modelling storm responses on a highenergy coastline with XBeach. Model. Earth Syst. Environ. 1:3.
 https://doi.org/10.1007/s40808-015-0003-8